

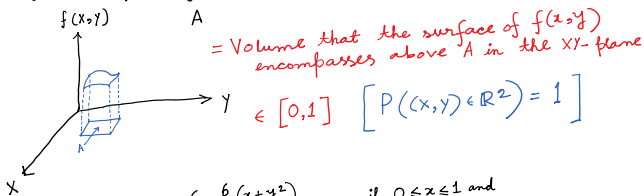
For two cont. r.v.'s X and Y ,

Joint pdf: $f(x,y)$.

- Requirements:
- $f(x,y) \geq 0 \quad \forall (x,y) \in \mathbb{R}^2$
 - $\iint_{\mathbb{R}^2} f(x,y) dy dx = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x,y) dy dx = 1$.

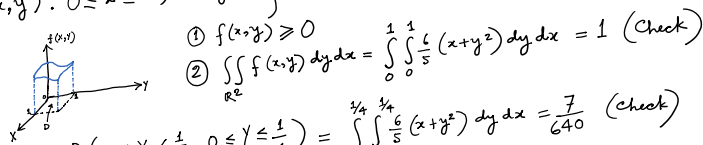
Joint Support: $D = \{(x,y) : f(x,y) > 0\}$.

$$P((X,Y) \in A) = \iint_A f(x,y) dy dx \quad [A \text{ is a set of points } (x,y)]$$



Eg 2: $f(x,y) = \begin{cases} \frac{6}{5}(x+y^2) & , \text{ if } 0 \leq x \leq 1 \text{ and } 0 \leq y \leq 1 \\ 0 & , \text{ o.w.} \end{cases}$

$$D = \{(x,y) : 0 \leq x \leq 1, 0 \leq y \leq 1\} = [0,1] \times [0,1] = [0,1]^2 : \text{Unit Square in } \mathbb{R}^2$$



- $f(x,y) \geq 0$
- $\iint_{\mathbb{R}^2} f(x,y) dy dx = \int_0^1 \int_0^1 \frac{6}{5}(x+y^2) dy dx = 1$ (check)

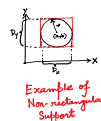
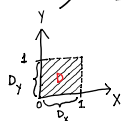
$$P(0 \leq X \leq \frac{1}{4}, 0 \leq Y \leq \frac{1}{4}) = \int_0^{\frac{1}{4}} \int_0^{\frac{1}{4}} \frac{6}{5}(x+y^2) dy dx = \frac{7}{640} \text{ (check)}$$

$$P(X = \frac{1}{4}, Y = \frac{1}{4}) = 0$$

$$P(X = \frac{1}{4}, 0 \leq Y \leq \frac{1}{4}) = 0$$

$$P(\frac{1}{4} \leq Y \leq \frac{3}{4}) = P(-\infty < X < \infty, \frac{1}{4} \leq Y \leq \frac{3}{4}) = \int_{-\infty}^{\infty} \int_{\frac{1}{4}}^{\frac{3}{4}} f(x,y) dy dx = \int_0^1 \int_{\frac{1}{4}}^{\frac{3}{4}} \frac{6}{5}(x+y^2) dy dx = \frac{37}{80} \text{ (check)}$$

Marginal Supports: $D_X = [0,1] = D_Y$
Here, $D = D_X \times D_Y$
(Rectangular Support)



$$\begin{cases} D_X \times D_Y = \{(x,y) : x \in D_X, y \in D_Y\} = [a-r, a+r] \times [b-r, b+r] \\ D = \{(x,y) : (x-a)^2 + (y-b)^2 \leq r^2\} : \text{Circle with center at } (a,b) \text{ and radius } r \end{cases}$$

(NOT rectangular)

Marginal pdf:

$$f_X(x) \stackrel{\text{defn.}}{=} \int_{y \in D_Y} f(x,y) dy$$

$$f_Y(y) \stackrel{\text{defn.}}{=} \int_{x \in D_X} f(x,y) dx$$

Eg 2. $f_X(x) = \begin{cases} \int_0^1 \frac{6}{5}(x+y^2) dy = \frac{6x}{5} + \frac{6}{5} \cdot \left[\frac{y^3}{3} \right]_0^1 = \frac{6x+2}{5} & , \text{ if } x \in [0,1] \\ 0 & , \text{ if } x \notin [0,1] \end{cases}$

$$f_X(x) \geq 0, \quad \int_0^1 f_X(x) dx = \int_0^1 \frac{6x+2}{5} dx = 1 \text{ (check)}$$

Similarly, $f_Y(y) = \begin{cases} \frac{6y^2+3}{5}, \text{ if } y \in [0,1] \\ 0, \text{ o.w.} \end{cases}$ (check)

Check that it's a density.
 $P(\frac{1}{4} \leq Y \leq \frac{3}{4}) = \int_{\frac{1}{4}}^{\frac{3}{4}} f_Y(y) dy = \int_{\frac{1}{4}}^{\frac{3}{4}} \frac{6y^2+3}{5} dy = \frac{37}{80}$ (check)

$$\left[\int_{-\infty}^{\infty} f_X(x) dx = \int_{-\infty}^{\infty} \left\{ \int_{-\infty}^{\infty} f(x,y) dy \right\} dx = 1 \right]$$

• Independence:

X, Y discrete: $A_x = \{X=x\}, B_y = \{Y=y\}$
D.I.M: X and Y are indep. if A_x and B_y are indep. $\forall (x,y)$.
i.e. $P(A_x \cap B_y) = P(A_x)P(B_y)$

equivalently, $P(A_x \cap B_y) = P(A_x) \cdot P(B_y)$

$$\Leftrightarrow f(x,y) = f_x(x) \cdot f_y(y), \forall (x,y)$$

X, Y conts: Indep. iff $f(x,y) = f_x(x) \cdot f_y(y), \forall (x,y)$

If X and Y are indep.,

$$P(a \leq X \leq b, c \leq Y \leq d) = P(a \leq X \leq b) \cdot P(c \leq Y \leq d)$$

$$\begin{aligned} \text{Ppf: LHS} &= \int_a^b \int_c^d f(x,y) dy dx = \int_a^b \int_c^d f_x(x) \cdot f_y(y) dy dx \\ &= \int_a^b f_x(x) \cdot \left[\int_c^d f_y(y) dy \right] dx \\ &= \left[\int_a^b f_x(x) dx \right] \left[\int_c^d f_y(y) dy \right] = \text{RHS} \end{aligned}$$

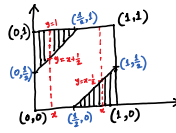
Extend for discrete case!

In general, $P(\underbrace{X \in A, Y \in B}_{P((X,Y) \in A \times B)}) = P(X \in A) \cdot P(Y \in B)$ if X and Y are indep.

Eg2: Same Joint density as yesterday.

$$P(|X-Y| > \frac{1}{2}) = P(X > Y + \frac{1}{2}) + P(Y > X + \frac{1}{2})$$

$$= \int_{x=0}^{\frac{1}{2}} \int_{y=x+\frac{1}{2}}^1 f(x,y) dy dx + \int_{x=\frac{1}{2}}^1 \int_{y=0}^{x-\frac{1}{2}} f(x,y) dy dx \text{ (finish)}$$



$Y = X + \frac{1}{2} \rightarrow$ Passes through $(0, \frac{1}{2})$ and $(\frac{1}{2}, 1)$
 $X = Y + \frac{1}{2} \rightarrow$ Passes through $(\frac{1}{2}, 0)$ and $(1, \frac{1}{2})$

Independence:

$p(x,y) = p_x(x) \cdot p_y(y), \forall (x,y)$ (equivalently $\forall (x,y) \in D_x \times D_y$)
 (OR, $f(x,y) = f_x(x) \cdot f_y(y), \forall (x,y)$ (equivalently $\forall (x,y) \in D_x \times D_y$)
 Entry in the table = (Corresponding Row Total) \times (Corresponding Col. Total)

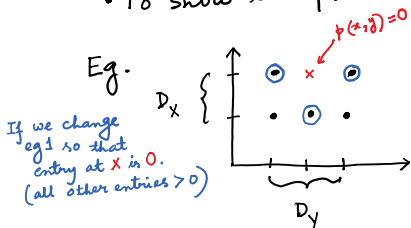
Let's look at Eg1, $p(100,0) = 0.2 \neq p_x(100) p_y(0)$

$p_x(100) = 0.5$
 $p_y(0) = 0.25$

$\Rightarrow X$ and Y are dep.

- To show dep, show $p(x,y) \neq p_x(x) \cdot p_y(y)$ for any one (x,y) .
- To show indep, show $p(x,y) = p_x(x) \cdot p_y(y)$ for all (x,y) .

Outside this rectangle both sides are 0.



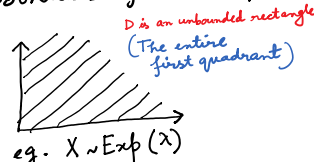
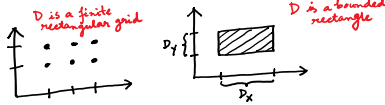
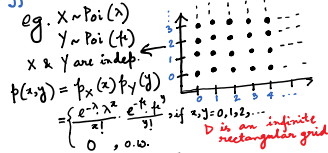
At x , $p(x,y) = 0 \neq p_x(x) \cdot p_y(y)$
 \Rightarrow NOT indep.
 ← These are true because of the \odot points.

Indep. \Rightarrow Rectangular Joint Support
 Non-rectangular Joint Support \Rightarrow Dep.

works for contin. case also.

If you have rectangular Joint Support, check pointwise for indep.

Rectangular Support:
 (Examples of different cases)



X and Y are indep.

Then $f(x,y) = f_x(x) \cdot f_y(y) = \begin{cases} \lambda e^{-\lambda x} \cdot \mu e^{-\mu y}, & \text{if } x,y > 0 \\ 0, & \text{o.w.} \end{cases}$

If I said: $f(x,y) \propto e^{-(\lambda x + \mu y)}, \forall (x,y) \text{ s.t. } x > 0, y > 0$, and then ask you to check for indep.

Ans: Joint Support is first Quadrant \rightarrow Rectangular (May or may not be indep.)
 Next $f(x,y) \propto e^{-\lambda x} e^{-\mu y} \rightarrow$ Indep. (Need not find out $f_x(x)$ and $f_y(y)$)

$f_1(x) = e^{-\lambda x}$, $f_2(y) = e^{-\mu y}$
 such that $\int_{-\infty}^{\infty} f_1(x) dx < \infty$ and $\int_{-\infty}^{\infty} f_2(y) dy < \infty$ } Just finding these two functions is enough for indep.

$f_x(x) = \int f(x,y) dy \propto \int f_1(x) f_2(y) dy = f_1(x) \cdot \int f_2(y) dy = K_1 f_1(x)$ (for some K_1)
 $\Rightarrow f_x(x) \propto f_1(x) \Rightarrow f_x(x) = K_1 f_1(x)$

Similarly, $f_y(y) = K_2 f_2(y)$ (for some K_2)

Now, $f(x,y) \propto f_1(x) f_2(y) \Rightarrow f(x,y) \propto f_x(x) f_y(y)$ [$f_1(x) f_2(y)$ and $f_x(x) f_y(y)$ differ by the constant $K_1 K_2$ only]
 ... $\int dx = 1$ (as f is a joint density) ... $\int dy = 1$ (as both $f_x(x)$ and $f_y(y)$ are densities)

Now, $f(x, y) \propto f_1(x) f_2(y) \Rightarrow f(x, y) \propto f_x(x) f_y(y)$ by the constant $K_1 K_2$ only
 But since $\iint f(x, y) dy dx = 1$ (as f is a joint density)
 and, $\int f_x(x) f_y(y) dy dx = (\int f_x(x) dx) (\int f_y(y) dy) = 1 \times 1 = 1$ (as both $f_x(x)$ and $f_y(y)$ are densities)
 $f(x, y) = f_x(x) \cdot f_y(y) = (K_1 f_1(x)) (K_2 f_2(y))$.

So if you can write $f(x, y)$ as a product of two functions $f_1(x)$ and $f_2(y)$, you have essentially proved independence (provided the support is rectangular) and you have also found out the marginal densities $f_x(x)$ and $f_y(y)$ without the proportionality constants.

In Eg 2., Joint Support = $[0, 1] \times [0, 1]$: Rectangular

But $f(x, y) = \frac{6}{5}(x+y^2)$ on $[0, 1] \times [0, 1]$

$$\neq \underbrace{\frac{6x+2}{5}}_{f_x(x)} \cdot \underbrace{\frac{6y^2+3}{5}}_{f_y(y)}$$

Check at $f(1, 1) = \frac{6}{5}(1+1^2) = \frac{12}{5}$

$f_x(1) = \frac{8}{5}, f_y(1) = \frac{9}{5}$

$\Rightarrow f(1, 1) \neq f_x(1) \cdot f_y(1) \Rightarrow$ NOT indep.

• Cond^t PMF:

$$P(A_x | B_y) = P(X=x | Y=y) = \frac{P(X=x, Y=y)}{P(Y=y)} = \frac{f(x, y)}{f_y(y)} \quad \left[\begin{array}{l} \text{Joint pmf} \\ \text{Marginal pmf} \end{array} \right. \left. \begin{array}{l} f_y(y) > 0 \\ \Leftrightarrow y \in D_y \end{array} \right]$$

Cond^t PMF of Y given $\{X=x\}$
 Note: $f_{Y|X}(y|x)$

$$\rightarrow P(Y=y | X=x) = \frac{f(x, y)}{f_x(x)} \quad [f_x(x) > 0 \Leftrightarrow x \in D_x]$$

• Cond^t PDF:

$$f_{Y|X}(y|x) = \frac{f(x, y)}{f_x(x)} \quad [x \in D_x]$$

$$f_{X|Y}(x|y) = \frac{f(x, y)}{f_y(y)} \quad [y \in D_y]$$

$p_{Y|X}(y|x)$ and $f_{Y|X}(y|x)$ give valid distributions when you consider all y -values $\in D_Y$. [For each $x \in D_X$]

Similar results for $p_{X|Y}(x|y)$ and $f_{X|Y}(x|y)$.

- $p_{Y|X}(y|x) \geq 0 \quad \forall x \in D_X$.
- $\sum_{y \in D_Y} p_{Y|X}(y|x) = \sum_{y \in D_Y} \frac{p(x,y)}{p_X(x)} = \frac{1}{p_X(x)} \cdot \sum_{y \in D_Y} p(x,y) = \frac{p_X(x)}{p_X(x)} = 1 \quad (\forall x \in D_X)$

Similar proof in conts. case and for cond. dist. of X given Y .

In Eg 1, $p_{X|Y}(100|0) = P(X=100|Y=0) = \frac{p(100,0)}{p_Y(0)} = \frac{0.2}{0.25} = \frac{4}{5}$
 $p_{X|Y}(250|0) = \frac{p(250,0)}{p_Y(0)} = \frac{0.05}{0.25} = \frac{1}{5}$
 $p_{X|Y}(x|0)$ is a valid pmf for X on $D_X = \{100, 250\}$

$p_X(100) = 0.5 = p_X(250)$
 NOT the same dist. as the cond. dist. of X given $Y=0$.
 \Rightarrow Dependent

Indep. $\Leftrightarrow p(x,y) = p_X(x) \cdot p_Y(y) \quad \forall (x,y)$
 $\Rightarrow p_X(x) = \frac{p(x,y)}{p_Y(y)} = p_{X|Y}(x|y) \quad \text{When } y \in D_Y$
 and, $p_Y(y) = p_{Y|X}(y|x) \quad \text{when } x \in D_X$

Indep. of two r.v.'s is equivalent to saying that the cond. dist. of one variable given the other variable is fixed at some value is same as the marginal dist. of the first variable.

• Cond. CDF: $F_{Y|X}(y|x) = \int_{-\infty}^y f_{Y|X}(z|x) dz$

• Cond. Mean: $E[Y|X=x] = \int_{-\infty}^{\infty} y f_{Y|X}(y|x) dy$
 $E[h(Y)|X=x] = \int_{-\infty}^{\infty} h(y) \cdot f_{Y|X}(y|x) dy$

• Cond. Percentile: $\eta_{Y|X}(p|x)$: Soln. of $F_{Y|X}(y|x) = p$ (Solve for y)
 Hence $F_{Y|X}(\eta_{Y|X}(p|x)|x) = p$.

Compare all these cond. quantities with corresponding uncond. quantities, i.e. the marginal quantities.
 eg. $F_Y(y), E(Y), E[h(Y)], \eta_Y(p), V(Y)$ etc.

• Cond. Var: $V(Y|X=x) = E[(Y - E(Y|X=x))^2 | X=x]$
 $= E[Y^2 | X=x] - E^2[Y|X=x] \quad (\text{check})$

Eg 2: Given $X=0.8$,
 $f_{Y|X}(y|0.8) = \frac{f(0.8,y)}{f_X(0.8)} = \frac{\frac{6}{5}(0.8+y^2)}{\frac{6}{5} \cdot 0.8 + \frac{2}{5}} \quad [f_X(x) = \frac{6x+2}{5} I_{[0,1]}(x)]$

$= \begin{cases} \frac{30y^2+24}{34}, & \text{if } y \in [0,1] \\ 0, & \text{o.w.} \end{cases}$

(This is a valid density on $[0,1] = D_Y$)
 - check

Compare it with $f_Y(y) = \frac{6y^2+3}{5} I_{[0,1]}(y) \quad (\Rightarrow X, Y \text{ are dep.})$

Compare it with $f_Y(y) = \frac{6y^2+3}{5} I_{[0,1]}(y)$ ($\Rightarrow \wedge, \vee, \dots$)

$$P(Y \leq 0.5 | X=0.8) = \int_0^{0.5} \frac{30y^2+24}{34} dy = 0.39 \text{ (check)}$$

$$E(Y | X=0.8) = \int_0^1 y \cdot \frac{30y^2+24}{34} dy = 0.574 \text{ (check)}$$

$\eta(0.5 | 0.8)$: Solve for: $F(y | 0.8) = 0.5$
 $y | x$

$$\int_0^y \frac{30z^2+24}{34} dz = 0.5 \Rightarrow 10y^3 + 24y = 17 \Rightarrow y = 0.6126$$

$$\Rightarrow \eta_{y|x}(0.5 | 0.8) = 0.6126.$$

More than 2 r.v.'s: X_1, \dots, X_n ($n > 2$)

(Discrete) Joint pmf: $p(x_1, \dots, x_n) = P(X_1=x_1, \dots, X_n=x_n) (\geq 0)$

$$\text{and } \sum_{x_1 \in D_{X_1}} \dots \sum_{x_n \in D_{X_n}} p(x_1, \dots, x_n) = 1.$$

(Conts.) Joint pdf: $f(x_1, \dots, x_n) \geq 0$

$$\text{and } \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f(x_1, \dots, x_n) dx_n \dots dx_1 = 1$$

$$\text{and, } P(a_1 \leq X_1 \leq b_1, \dots, a_n \leq X_n \leq b_n) = \int_{a_1}^{b_1} \dots \int_{a_n}^{b_n} f(x_1, \dots, x_n) dx_n \dots dx_1$$

X_1, X_2, \dots, X_n are indep. if:

For every subset $X_{i_1}, X_{i_2}, \dots, X_{i_k}$ of size k ($2 \leq k \leq n$) of X_1, \dots, X_n :

$$p(x_{i_1}, \dots, x_{i_k}) = p_{X_{i_1}}(x_{i_1}) \dots p_{X_{i_k}}(x_{i_k}) \text{ (Discrete)}$$

$$\text{or, } f(x_{i_1}, \dots, x_{i_k}) = f_{X_{i_1}}(x_{i_1}) \dots f_{X_{i_k}}(x_{i_k}) \text{ (Conts.)}$$

Sec 5.2: If $h(x, y)$ is a 2-dim fn.

$$E[h(x, y)] = \begin{cases} \sum_{x \in D_X} \sum_{y \in D_Y} h(x, y) p(x, y). & \text{(Discrete)} \\ \int \int h(x, y) f(x, y) dy dx. & \text{(Conts.)} \end{cases}$$

If $h(x, y) = h(x)$,

$$E[h(x, y)] = \sum_{x \in D_X} \sum_{y \in D_Y} h(x) \cdot p(x, y)$$

$$= \sum_{x \in D_X} h(x) \sum_{y \in D_Y} p(x, y)$$

$$= \sum_{x \in D_X} h(x) \cdot p_X(x) = E[h(x)]$$

Eg 1: $h(x, y) = |x-y|$

$h(x, y) | \dots \dots 200$

$$E[h(x, y)] = E|x-y|$$

$$= 100 \times 0.2 + 0 \times 0.1 + 100 \times 0.2$$

Eg 1: $h(x,y) = |x-y|$

$h(x,y)$ $x \backslash y$	0	100	200
100	100	0	100
250	250	150	50

$$E[h(x,y)] = E|x-y|$$

$$= 100 \times 0.2 + 0 \times 0.1 + 100 \times 0.2 + 250 \times 0.05 + 150 \times 0.15 + 50 \times 0.3 = 90 \text{ (check)}$$

Defn: $V[h(x,y)] = E \left[\underbrace{\{h(x,y) - E[h(x,y)]\}}_{k(x,y)} \right]^2$

$$= E[h^2(x,y)] - E^2[h(x,y)] \text{ (check)}$$

Also, if $g(x,y) = h(x,y) + k(x,y)$,

$$E[g(x,y)] = E[h(x,y)] + E[k(x,y)] \text{ (check)}$$

Eg 2: $h(x,y) = \frac{x}{2} + y + \frac{1}{2}$

$$E[h(x,y)] = E\left[\frac{x}{2}\right] + E[y] + \frac{1}{2}$$

$$= \frac{1}{2} \cdot E[x] + E[y] + \frac{1}{2}$$

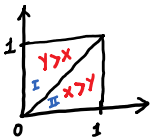
$$= \frac{1}{2} \cdot \frac{3}{5} + \frac{3}{5} + \frac{1}{2} \text{ (check)}$$

$$= 1.4$$

Lecture 28

Monday, July 24, 2017 1:01 PM

Eg 2:



$$\begin{aligned}
 E|X-Y| &= \int_0^1 \int_0^1 |x-y| \frac{6}{5} (x+y^2) dy dx \\
 &= \int_I |x-y| \frac{6}{5} (x+y^2) dy dx + \int_{II} |y-x| \frac{6}{5} (x+y^2) dy dx \\
 &= \int_{x=0}^1 \int_{y=x}^1 (y-x) \frac{6}{5} (x+y^2) dy dx + \int_{x=0}^1 \int_{y=0}^x (x-y) \frac{6}{5} (x+y^2) dy dx
 \end{aligned}$$

(finish)

In general, to calculate $E|ax+by+c|$ or $P(|ax+by+c| > 0)$, look at 2 sides of the straight line $ax+by+c=0$.

• Covariance: (b/w 2 r.v.'s X and Y)

$$\text{Cov}(X, Y) \stackrel{\text{def}}{=} E \left[\underbrace{(X - \mu_X)(Y - \mu_Y)}_{h(X, Y)} \right]$$

$$= \begin{cases} \sum_{x \in D_X} \sum_{y \in D_Y} (x - \mu_X)(y - \mu_Y) p(x, y) & (\text{Discrete}) \\ \iint_{x \in D_X} \int_{y \in D_Y} (x - \mu_X)(y - \mu_Y) f(x, y) dy dx & (\text{Conts.}) \end{cases}$$

$$h(x, y) = (x - \mu_X)(y - \mu_Y) > 0, \text{ if } \begin{cases} x > \mu_X \text{ and } y > \mu_Y \\ \text{or } x < \mu_X \text{ and } y < \mu_Y \end{cases}$$

$$< 0, \text{ o.w.}$$

$$\Rightarrow \text{Cov}(X, Y) > 0 \text{ if } \begin{cases} X \uparrow, Y \uparrow \\ X \downarrow, Y \downarrow \end{cases}$$

$$< 0 \text{ if } \begin{cases} X \uparrow, Y \downarrow \\ X \downarrow, Y \uparrow \end{cases}$$

Eg 1. $E(X) = 100 \times 0.5 + 250 \times 0.5 = 175$

$E(Y) = 125$ (check)

$$\text{Cov}(X, Y) = (100 - 175)(0 - 125) \times 0.2 + \dots + (250 - 175)(200 - 125) \times 0.3 = 1875 \text{ (check)}$$

Properties:

- $\text{Cov}(X, Y) = \text{Cov}(Y, X)$
- $\text{Cov}(X, Y) = E(XY) - E(X) \cdot E(Y)$

Pf: $\text{Cov}(X, Y) = E[(X - \mu_X)(Y - \mu_Y)]$

$$\begin{aligned}
 &= E[XY - \mu_X Y - \mu_Y X + \mu_X \mu_Y] \\
 &= E(XY) - \mu_X E(Y) - \mu_Y E(X) + \mu_X \mu_Y \\
 &= E(XY) - E(X) \cdot E(Y)
 \end{aligned}$$

- $\text{Cov}(X, X) = E[(X - \mu_X)^2] = \text{Var}(X)$

$$3. \text{Cov}(X, X) = E[(X - \mu_X)^2] = \text{Var}(X).$$

$$\begin{aligned} \text{[Eg 2: } E(X) = E(Y) = \frac{3}{5} \\ E(XY) = \int_0^1 \int_0^1 xy \cdot \frac{6}{5}(x+y^2) dy dx = \frac{6}{5} \left(\int_0^1 x^2 dx \right) \left(\int_0^1 y^2 dy \right) + \frac{6}{5} \left(\int_0^1 x dx \right) \left(\int_0^1 y^3 dy \right) \\ = \frac{7}{20} \text{ (check)} \end{aligned}$$

$$\text{Cov}(X, Y) = \frac{7}{20} - \frac{9}{25} = -0.01$$

$$4. \text{Cov}(aX+b, Y) = a \text{Cov}(X, Y)$$

$$\text{Pf: } \text{Cov}(aX+b, Y) = E[(aX+b - a\mu_X - b)(Y - \mu_Y)] \\ = a \cdot \text{Cov}(X, Y).$$

$$5. \text{Cov}(aX+b, cY+d) = ac \text{Cov}(X, Y)$$

Apply 4 for both variables.

If X & Y are measured in Unit 1 and Unit 2, then $\text{Cov}(X, Y)$ has unit Unit 1 \times Unit 2.

Correlation:

$$\text{Corr}(X, Y) \stackrel{\text{Notn}}{=} \rho_{X, Y} \stackrel{\text{Defn}}{=} \frac{\text{Cov}(X, Y)}{\sqrt{\text{V}(X) \cdot \text{V}(Y)}} : \text{Unit free [as } \text{V}(X) \text{ has unit (Unit 1)}^2 \text{ and } \text{V}(Y) \text{ has unit (Unit 2)}^2 \text{.]}$$

Properties: 1. $\rho_{X, Y} = \rho_{Y, X}$

$$2. \rho_{X, X} = \frac{\text{Cov}(X, X)}{\sqrt{[\text{V}(X)]^2}} = \frac{\text{V}(X)}{\text{V}(X)} = 1.$$

$$3. \boxed{-1 \leq \rho_{X, Y} \leq 1}$$

(i) $\rho_{X, Y} = 0$: Uncorrelated (NO linear association)

(ii) $|\rho_{X, Y}| = 1$: Perfect linear association

(iii) $|\rho_{X, Y}| \geq 0.8$: Strong association

(iv) $0.5 < |\rho_{X, Y}| < 0.8$: Moderate association

(v) $|\rho_{X, Y}| \leq 0.5$: Weak association

$|\rho_{X, Y}|$: Strength of the association

Sign of $\rho_{X, Y}$: Direction

$$\begin{aligned} \text{[Eg 1. } \text{Cov}(X, Y) = 1875 \\ \text{V}(X) = 5625, \text{V}(Y) = 6875 \text{ (check)} \\ \rho_{X, Y} = 0.301 \text{ (check)} \end{aligned}$$

Lecture 29

Tuesday, July 25, 2017 1:00 PM

④ $|\rho_{ax+b, cy+d}| = |\rho_{x,y}|$

Pf. $\text{Cov}(ax+b, cy+d) = ac \cdot \text{Cov}(x,y)$

$V(ax+b) = a^2 \cdot V(x)$

$V(cy+d) = c^2 \cdot V(y)$

$\rho_{ax+b, cy+d} = \frac{\text{Cov}(ax+b, cy+d)}{\sqrt{V(ax+b) \cdot V(cy+d)}}$

$= \frac{ac \cdot \text{Cov}(x,y)}{\sqrt{a^2 V(x) \cdot c^2 V(y)}} = \frac{ac \cdot \text{Cov}(x,y)}{|a| |c| \sqrt{V(x) \cdot V(y)}}$

$= \frac{ac}{|a| \cdot |c|} \cdot \rho_{x,y}$

$|\rho_{ax+b, cy+d}| = |\rho_{x,y}|$

Moreover, $\rho_{ax+b, cy+d} = \pm \rho_{x,y}$ [$+$ if a and c have the same sign, $-$ if a and c have opposite signs]

⑤ $\rho_{ax+b, x} = \pm \rho_{x,x} = \pm 1$ ($+$ if $a > 0$, $-$ if $a < 0$)

Use 4 with $y=x$ and $c=1$ and $d=0$.

Sp. case: $\rho_{x, -x} = -1 = \rho_{-2x+3, x}$

Q: What if $a=0$? (If $a=0$, $ax+b=b$ w.p.1)

$\rho_{ax+b, x} = \rho_{b, x} = \frac{\text{Cov}(x, b)}{\sqrt{\sigma_x^2 \cdot \sigma_b^2}}$

$= \frac{0}{\sigma_x \cdot 0}$

: NOT defined

$\left[\begin{matrix} \sigma_b^2 = 0 \\ \text{Cov}(x, b) = E[(x-\mu_x)(b-b)] = 0 \end{matrix} \right]$

Fact: $\rho_{x,y}$ is NOT defined if X or Y is degenerate.

Eg 2: $\text{Cov}(x,y) = -0.01$

$E(x) = E(y) = \frac{3}{5}$

$V(x) = \frac{11}{150}, V(y) = \frac{2}{250}$ (Check)

$\rho_{x,y} = \frac{-0.01}{\sqrt{\frac{11}{150} \times \frac{2}{250}}} = -0.131$

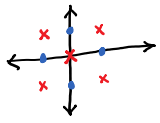
• Comparison of Indep. & Uncorrelatedness:

a. Indep. \Rightarrow Uncorrelated

b. Uncorrelated $\not\Rightarrow$ Indep.

a. Indep. $\Rightarrow E(XY) = E(X) \cdot E(Y) \Rightarrow \text{Cov}(X, Y) = 0 \Rightarrow (X, Y) = U.$

b. Counter Example:



$$D = \{(-1, 0), (0, 1), (0, -1), (1, 0)\}$$

$$p(x, y) = \begin{cases} 0.25, & \text{if } (x, y) \in D \\ 0, & \text{o.w.} \end{cases}$$

$$D_x = \{-1, 0, 1\} = D_y$$

Non-rectangular support \Rightarrow NOT indep. (Look at any x)

$$p(x, y) = 0 \neq p_x(x) \cdot p_y(y)$$

\downarrow \downarrow
 0 0
 e.g. $p(0, 0) = 0 \neq p_x(0) \cdot p_y(0) = 0.5 \times 0.5$

$$XY = 0 \text{ w.p. } 1 \Rightarrow E(XY) = 0$$

$$\text{Also, } p_x(x) = \begin{cases} .25, & \text{if } x = -1 \\ .5, & \text{if } x = 0 \\ .25, & \text{if } x = 1 \end{cases} \Rightarrow E(X) = 0$$

$$\text{Similarly, } E(Y) = 0.$$

$$\Rightarrow \text{Cov}(X, Y) = 0 - 0 \cdot 0 = 0 \Rightarrow \rho_{X, Y} = 0. \text{ (Clearly, } \sigma_X, \sigma_Y > 0)$$

Correlation $\not\Rightarrow$ Causation

• Uniform Distⁿ for (X, Y) : (cont.)

$$f(x, y) = K I_A(x, y), \quad D = A$$

$$\iint_{\mathbb{R}^2} f(x, y) dy dx = \iint_A K \cdot dy dx = K \iint_A 1 \cdot dy dx = K |A| = 1$$

$$\Rightarrow K = \frac{1}{|A|}$$

$$P(B) = \begin{cases} \frac{|B|}{|A|}, & \text{if } B \subset A \\ \frac{|B \cap A|}{|A|}, & \text{in general} \end{cases}$$

$$[|A| = \text{area}(A)]$$

• Sec 5.5: A linear combination of X_1, \dots, X_n (RV's) with means

$\mu_1, \mu_2, \dots, \mu_n$ and SD's $\sigma_1, \sigma_2, \dots, \sigma_n$ is defined as:

$$a_1 X_1 + a_2 X_2 + \dots + a_n X_n \text{ for constants } a_1, \dots, a_n.$$

$$1. E(a_1 X_1 + \dots + a_n X_n) = a_1 E(X_1) + \dots + a_n E(X_n) = a_1 \mu_1 + \dots + a_n \mu_n$$

$$2. V(a_1 X_1 + \dots + a_n X_n) = a_1^2 V(X_1) + \dots + a_n^2 V(X_n) = a_1^2 \sigma_1^2 + \dots + a_n^2 \sigma_n^2 \text{ [If } X_1, \dots, X_n \text{ are indep.]}$$

$$3. V(a_1 X_1 + \dots + a_n X_n) = \sum_{i=1}^n \sum_{j=1}^n a_i a_j \text{Cov}(X_i, X_j)$$

4. If $X_i \sim N(\mu_i, \sigma_i^2)$ are indep,

$$a_1 X_1 + \dots + a_n X_n \sim N(a_1 \mu_1 + \dots + a_n \mu_n, a_1^2 \sigma_1^2 + \dots + a_n^2 \sigma_n^2).$$

e.g. $E(X_1 - X_2) = E(X_1) - E(X_2)$ } Take $n=2, a_1=1, a_2=-1$
 $V(X_1 - X_2) = V(X_1) + V(X_2)$ } in 1 and 2

$$a_1 X_1 + \dots + a_n X_n \sim N(\dots)$$

e.g. $E(X_1 - X_2) = E(X_1) - E(X_2)$ } Take $n=2, a_1=1, a_2=-1$
in 1 and 2
 $V(X_1 - X_2) = V(X_1) + V(X_2)$
(if X_1 and X_2 are indep.)

Moreover if $X_1 \sim N(\mu_1, \sigma_1^2)$ and $X_2 \sim N(\mu_2, \sigma_2^2)$ are indep.,
 $X_1 - X_2 \sim N(\mu_1 - \mu_2, \sigma_1^2 + \sigma_2^2)$ (By 4)

Sketches of Proofs:

$$1. \text{ For } n=2, E(a_1 X_1 + a_2 X_2) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (a_1 x_1 + a_2 x_2) f(x_1, x_2) dx_2 dx_1$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} a_1 x_1 f(x_1, x_2) dx_2 dx_1 + \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} a_2 x_2 f(x_1, x_2) dx_2 dx_1$$

$$= a_1 \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x_1 f(x_1, x_2) dx_2 dx_1 + a_2 \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x_2 f(x_1, x_2) dx_2 dx_1$$

$$= a_1 E(X_1) + a_2 E(X_2)$$

} Similar proof in the discrete case.

For $n=3, E(a_1 X_1 + a_2 X_2 + a_3 X_3) = E(Y + a_3 X_3) = E(Y) + E(a_3 X_3)$ [By using result for $n=2$ with $a_1=1, X_1=Y, a_2=a_3, X_2=X_3$]
(Let $Y = a_1 X_1 + a_2 X_2$) $= E(a_1 X_1 + a_2 X_2) + a_3 E(X_3)$
 $= a_1 E(X_1) + a_2 E(X_2) + a_3 E(X_3)$ [Using result for $n=2$ again]

Proof by induction.

$$3. V(aX + bY) = E[(aX + bY)^2] - [E(aX + bY)]^2$$

$$= E[a^2 X^2 + b^2 Y^2 + 2abXY] - [aE(X) + bE(Y)]^2$$

$$= a^2 E(X^2) + b^2 E(Y^2) + 2ab E(XY) - a^2 E^2(X) - b^2 E^2(Y) - 2ab E(X) \cdot E(Y)$$

$$= a^2 \{E(X^2) - E^2(X)\} + b^2 \{E(Y^2) - E^2(Y)\} + 2ab \{E(XY) - E(X) \cdot E(Y)\}$$

$$= a^2 V(X) + b^2 V(Y) + \frac{2ab \text{Cov}(X, Y)}{ab \text{Cov}(X, Y) + ba \text{Cov}(Y, X)}$$

Similar proof by induction.

2. If X_1, \dots, X_n are indep., $\text{Cov}(X_i, X_j) = 0 \quad \forall i \neq j$
By 3, $V(a_1 X_1 + a_2 X_2 + \dots + a_n X_n) = \sum_{i=1}^n \sum_{j=1}^n a_i a_j \text{Cov}(X_i, X_j) = \sum_{i=1}^n a_i^2 \text{Cov}(X_i, X_i) = \sum_{i=1}^n a_i^2 V(X_i)$

Proof of 4 is beyond the scope of this book.